Equity Premia as Low as Three Percent?
Evidence from Analysts’ Earnings Forecasts
for Domestic and International Stock Markets

JAMES CLAUS and JACOB THOMAS*

ABSTRACT
The returns earned by U.S. equities since 1926 exceed estimates derived from
theory, from other periods and markets, and from surveys of institutional inves-
tors. Rather than examine historic experience, we estimate the equity premium
from the discount rate that equates market valuations with prevailing expecta-
tions of future flows. The accounting flows we project are isomorphic to projected
dividends but use more available information and narrow the range of reasonable
growth rates. For each year between 1985 and 1998, we find that the equity pre-
mium is around three percent (or less) in the United States and five other markets.

The equity risk premium lies at the core of financial economics. Representing
the excess of the expected return on the stock market over the risk-free rate,
the equity premium is unobservable and has been estimated using different
approaches and samples. The estimates most commonly cited in the aca-
demic literature are from Ibbotson Associates’ annual reviews of the per-
formance of various portfolios of U.S. stocks and bonds since 1926. Those
estimates lie in the region of seven to nine percent per year, depending on
the specific series examined. This historic evidence is objective and easy to
interpret and has convinced many, especially academic financial economists,
that the Ibbotson estimates are the best available proxies for the equity
premium (Welch (1999)). For discussion purposes, we use “eight percent”

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Jimmy Liew, Jing Liu, Jim McKeown, Karl Muller, Jim Ohlson, Stephen Penman, Huai Zhang,
and workshop participants at AAA annual meetings (San Diego), Columbia University, Copen-
hagen Business School, University of Michigan, Michigan State University, University of North
Carolina-Chapel Hill, Northern Arizona University, Ohio State University, Penn State Univer-
sity, Prudential Securities Quantitative Conference, Syracuse University, and University of
Texas-Austin.

1 The annualized distribution of monthly common stock returns over the 30-day T-bill rate
has a mean of 9.12 percent and a standard deviation of 20.06 percent (from data in Table A-16,
Ibbotson Associates (1999)). If these 73 observations are independent and identically distrib-
uted, the sample mean is a reasonable estimate for the equity premium, and the standard error
of 2.35 percent associated with the sample mean allows an evaluation of other hypothesized
values of the equity premium.
and “the Ibbotson estimate” interchangeably to represent the historic mean of excess returns earned by U.S. equities since 1926. (Unless noted otherwise, all amounts and rates are stated in nominal, not real, terms.)

Our objective is to show empirically that eight percent is too high an estimate for the equity premium in recent years. Rather than examine observed returns, we estimate for each year since 1985 the discount rate that equates U.S. stock market valuations with the present value of prevailing forecasts of future flows. Subtracting 10-year risk-free rates from these estimated discount rates suggests that the equity premium is only about three percent. An examination of five other large stock markets (Canada, France, Germany, Japan, and the United Kingdom) provides similar results. Despite substantial variation in the underlying fundamentals across markets and over time, observing that every one of our 69 country-year estimates lies well below eight percent suggests that the Ibbotson estimate is too high for our sample period. Examination of various diagnostics (such as implied future profitability) confirms that the projections required to support an eight percent equity premium are unreasonable and inconsistent with past experience.

Some features of our study should be emphasized at the outset. As we only seek to establish a reasonable upper bound for the equity premium, we select long-term growth assumptions that exceed past experience and do not adjust for optimism in the analyst forecasts used. Also, we use the simplest structure necessary to conduct our analysis. Our estimates refer to a long-term premium expected to hold over all future years (whereas historical estimates measure one-period premia), and we assume that the premium is constant over those future years (we do incorporate anticipated variation in risk-free rates). Finally, each annual estimate is conditional on the information available in that year; we do not consider an unconditional equity premium toward which those conditional premia might gravitate in the long run.

We are not the first to question the validity of the Ibbotson estimate. Mehra and Prescott (1985) initiated a body of theoretical work that has examined the so-called “equity premium puzzle.” Their model indicates that the variance–covariance matrix of aggregate consumption and returns on stocks and bonds, when combined with reasonable risk-aversion parameters, implies equity premium estimates that are less than one percent. Despite subsequent efforts to bridge this gap (e.g., Abel (1999)), concerns remain about the validity of the Ibbotson estimate (see Kocherlakota (1996), Cochrane (1997), and Siegel and Thaler (1997) for summaries).

Gebhardt, Lee, and Swaminathan (forthcoming) find similar results when estimating firm-specific discount rates, rather than the market-level discount rates considered in this paper.

As described later, analyst optimism has declined systematically over time and a simple adjustment for mean bias is inappropriate. Bayesian adjustments to control for observed analyst optimism are not considered because we focus on an upper bound. In general, we do not use more complex econometric techniques and data refinements that are available to get sharper point estimates (e.g., Mayfield (1999), Vuolteenaho (1999), and Ang and Liu (2000)).
Surveys of institutional investors also suggest an equity premium substantially below eight percent (e.g., Burr (1998)), and there are indications that this belief has been held for many years (e.g., Benore (1983)). Also, the weighted average cost of capital used in discounted cash flow valuations provided in analysts' research reports usually implies an equity premium below five percent. Current share prices appear systematically overpriced if an eight percent equity premium is used on reasonable projections of future flows. This overpricing is more evident when examining mature firms, where there is less potential for disagreement about growth opportunities.

To identify possible reasons why the Ibbotson estimate might overstate the equity premium in recent years, apply the Campbell (1991) decomposition of observed returns in excess of the expected risk-free rate for the market portfolio. The four components are: (1) the expected equity premium for that period; (2) news about the equity premium for future periods; (3) news about current and future period real dividend growth; and (4) news about the real risk-free rate for current and future periods. Here, news represents changes in expectations between the beginning and end of the current period (for current period dividend growth and risk-free rates, it represents the unexpected portion of observed values). Summing up both sides of this relation for each year since 1926 indicates that the average excess return observed would exceed the equity premium today if: (1) conditional one-year-ahead equity premia have declined; (2) the conditional long-term equity premium anticipated for future years has declined; (3) news about real dividend growth was positive on average; or (4) the expected real risk-free rate has declined.

The first and second reasons for why the Ibbotson estimate overstates the current equity premium highlight the potential pitfalls of estimating equity premia from observed returns. Holding aside news about dividends and risk-free rates, valuations would exceed expectations if the equity premium has declined (since present values increase when expected rates of return decline). That is, unexpected changes in the equity premium cause historical equity premium estimates to move in the opposite direction. Blanchard (1993) concludes that the equity premium has declined since 1926 to two or three percent by the early 1990s, and speculates that this decline is caused by a simultaneous decline in expected real rates of return on stocks and an increase in expected real risk-free rates. (This increase in expected real risk-free rates is another puzzle, but that puzzle is beyond the scope of this paper.) The remarkable run-up in stock prices during the 1990s, both domestically as well as internationally, is also consistent with a recent decline

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4 While many argue for an equity premium between two and three percent (e.g., Bogle (1999, p. 76)), some suggest that the premium is currently close to zero (e.g., Glassman and Hassett (1998), and Wien (1998)). Surveys of individual investors, on the other hand, suggest equity premia even higher than the Ibbotson estimate. For example, the New York Times (October 10, 1997, page 1, “High hopes of mutual fund investors”), reported an equity premium in excess of 16 percent from a telephone survey conducted by Montgomery Asset Management.
in the equity premium. Stulz (1999) argues that increased globalization has caused equity premia to decline in all markets.

Examination of historic evidence over other periods and markets suggests that the U.S. experience since 1926 is unusual. Siegel (1992) finds that the excess of observed annual returns for NYSE stocks over short-term government bonds is 0.6, 3.5, and 5.9 percent over the periods 1802 to 1870, 1871 to 1925, and 1926 to 1990, respectively. Jorion and Goetzmann (1999) examine the evidence for 39 equity markets going back to the 1920s, and conclude that the high equity premium observed in the United States appears to be the exception rather than the rule. Perhaps some stock markets collapsed and those markets that survived, like the U.S. exchanges, exhibit better performance than expected (see Brown, Goetzmann, and Ross (1995)). This evidence is consistent with the third reason for the high Ibbotson premium: since 1926, news about real dividend growth for U.S. stocks has been positive on average.

Partially in response to these limitations of inferring equity premia from observed returns, financial economists have considered forward-looking approaches based on projected dividends.\(^5\) Informally, expected rates of return on the market equal the forward dividend yield plus expected growth in dividends (this dividend growth model is discussed in Section I). While dividend yields are easily measured, expected dividend growth in perpetuity is harder to identify. Proxies used for expected dividend growth include observed growth in earnings, dividends, or economy-wide aggregates (e.g., Fama and French (2000)). Unfortunately, the dividend growth rate that can be sustained in perpetuity is a hypothetical rate that is not necessarily anchored in any observable series, leaving considerable room for disagreement (see the Appendix for explanation).

We use a different forward-looking approach, labeled the abnormal earnings (or residual income) model, to mitigate problems associated with the dividend growth model.\(^6\) Recognizing that dividends equal earnings less changes in accounting (or book) values of equity allows the stream of projected dividends to be replaced by the current book value of equity plus a function of future accounting earnings (details follow in Section I). While book values feature prominently in the model, the inclusion of future abnormal earnings makes it isomorphic to the dividend discount model. Relative to the dividend growth model, this approach makes better use of currently

\(^5\) A related approach is to run predictive regressions of market returns or equity premium on dividend yields and other variables (e.g., Campbell and Shiller (1988)). We do not consider that approach because the declining dividend yields in recent years have caused predicted equity premium to turn negative (e.g., Welch (1999)).

\(^6\) The approach appears to have been discovered independently by a number of economists and accountants over the years. Preinreich (1938) and Edwards and Bell (1961) are two early cites. More recently, a large body of analytical and empirical work has utilized this insight (e.g., Penman (1999)). Examples of empirical investigations include market myopia (Abarbanell and Bernard (1999)), explaining cross-sectional variation in returns (Liu and Thomas (2000)), and stock picking (Frankel and Lee (1998a, 1998b)).
available information to reduce the importance of assumed growth rates, and it narrows the range of allowable growth rates by focusing on growth in rents, rather than dividend growth.

If the equity premium is as low as our estimates suggest, required rates of return (used for capital budgeting, regulated industries, and investment decisions) based on the Ibbotson estimate are severely overstated. Second, a smaller equity premium reduces the importance of estimating beta accurately (because required rates of return become less sensitive to variation in beta) and increases the magnitude of beta changes required to explain abnormal returns observed for certain market anomalies. Finally, reducing substantially the magnitude of the equity premium puzzle to be explained might reinvigorate theory-based studies.

In Section I we develop the abnormal earnings approach used in this paper and compare it with the dividend growth model. Section II contains a description of the sample and methodology. The equity premium estimates for the United States are reported in Section III, and those for the five other markets are provided in Section IV. To confirm that our estimates are robust, we conducted extensive sensitivity analyses, which we believe represent an important contribution of our research effort. A summary of that investigation is reported in Section V (details are provided in Claus and Thomas (1999a)) and Section VI concludes.

I. Dividend Growth and Abnormal Earnings Models

The Gordon (1962) dividend growth model is described in equation (1). This relation implies that the expected rate of return on the stock market ($k^*$) equals the forward dividend yield ($d_1/p_0$) plus the dividend growth rate in perpetuity ($g$) expected for the market.

$$p_0 = \frac{d_1}{k^* - g} \Rightarrow k^* = \frac{d_1}{p_0} + g$$  \hspace{1cm} (1)

where

- $p_0 =$ current price, at the end of year 0,
- $d_t =$ dividends expected at the end of future year $t$,
- $k^*$ = expected rate of return on the market, derived from the dividend growth model, and
- $g =$ expected dividend growth rate, in perpetuity.

The Gordon growth model is a special case of the general Williams (1938) dividend discount model, detailed in equation (2), where dividend growth is constrained to equal $g$ each year.

$$p_0 = \frac{d_1}{(1 + k^*)} + \frac{d_2}{(1 + k^*)^2} + \frac{d_3}{(1 + k^*)^3} + \ldots$$  \hspace{1cm} (2)
Research using the dividend growth model has often assumed that $g$ equals forecasted earnings growth rates obtained from sell-side equity analysts, who provide earnings forecasts along with their buy/sell recommendations. These forecasts refer to earnings growth over the next “cycle,” which is commonly interpreted to represent the next five years. Consequently, we refer to this earnings growth forecast as $g_5$. While most studies using $g_5$ as a proxy for $g$ have focused on the U.S. market alone (e.g., Brigham, Shome, and Vinson (1985)), some have examined other major equity markets also (e.g., Khorana, Moyer, and Patel (1997)). Estimates of the equity premium based on the assumption that $g$ equals $g_5$ are similar in magnitude to the Ibbotson estimate derived from historical data. For example, Moyer and Patel (1997) estimate the equity premium each year over their 11-year sample period (1985 to 1995) and generate a mean estimate of 9.38 (6.96) percent relative to the 1-year (30-year) risk-free rate.

However, others have balked at using $g_5$ as a proxy for $g$ (e.g., Malkiel (1996), Cornell (1999)) because it appears unreasonably high at an intuitive level, and have stepped down assumed growth rates. Forecasted values of $g_5$ for the United States over our sample period, which are close to 12 percent in all years, exceed nominal growth in S&P earnings, which has been only 6.6 percent since the 1920s (Wall Street Journal, June 16, 1997, “As stocks trample price measures, analysts stretch to justify buying”). Also, the real growth rate implied by the nominal 12 percent earnings growth rate exceeds both forecast and realized growth in GDP (since 1970, forecasts of expected real growth in GDP have averaged 2.71 percent, and realized real growth has averaged 2.81 percent).

While we show that $g_5$ is systematically optimistic relative to realized earnings, it is difficult to infer reliably the level of that optimism from the relatively short time-series of forecast errors available (reliable data on analyst forecasts go back only about 15 years). Moreover, the incentives for analysts to make optimistic forecasts vary across firms and over time. For example, the literature on U.S. analysts’ forecasts suggests that while analysts tended to make optimistic forecasts early in our sample period (to curry favor with management), more recently, management has tended to guide near-term analyst forecasts downward to be able to meet or beat them when announcing earnings. Even if unbiased estimates of near-term earnings growth ($g_5$) were available, the Appendix describes why those estimates as well as observed growth rates are conceptually different from $g$, the hypothetical dividend growth that can be sustained in perpetuity.

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7 Results reported in Table VI offer clear evidence of such a decline in optimism for all horizons. Bagnoli, Beneish, and Watts (1999) document how recent analyst forecasts are systematically below reported earnings for their sample, and also below “whisper” forecasts that are generally viewed as representing the market’s true earnings expectations. Matsumoto (1999) offers evidence in support of management guiding analyst forecasts downward, and also investigates factors that explain cross-sectional variation in this propensity to guide analysts.
The abnormal earnings model is an alternative that mitigates many of the problems noted above. Expected dividends can be related to forecasted earnings using equation (3) below, and that relation allows a conversion of the discounted dividends relation in equation (2) to the abnormal earnings relation in equation (4).

\[ d_t = e_t - (bv_t - bv_{t-1}) \]  

\[ p_0 = bv_0 + \frac{ae_1}{(1 + k)} + \frac{ae_2}{(1 + k)^2} + \frac{ae_3}{(1 + k)^3} + \ldots, \]  

where

- \( e_t \) = earnings forecast for year \( t \),
- \( bv_t \) = expected book (or accounting) value of equity at the end of year \( t \),
- \( ae_t = e_t - k(bv_{t-1}) \) = expected abnormal earnings for year \( t \), or forecast accounting earnings less a charge for the cost of equity, and
- \( k \) = expected rate of return on the market portfolio, derived from the abnormal earnings model.

Equation (3), also known as the “clean surplus” relation, requires that all items affecting the book value of equity (other than transactions with shareholders, such as dividends and share repurchases/issues) be included in earnings. Under U.S. accounting rules, almost all transactions satisfy the clean-surplus assumption. An examination of the few transactions that do not satisfy this relation suggests that these violations occur ex post, and are not anticipated in analysts’ earnings forecasts (e.g., Frankel and Lee (1998b)). Since we construct future book values using equation (3), by adding forecast income to and subtracting forecast dividends from beginning book values, clean surplus is maintained and the dividend and abnormal earnings relations in equations (2) and (4) are isomorphic.

Equation (4) shows that the current stock price equals the current book value of equity plus the present value of future expected abnormal earnings. Abnormal earnings, a proxy for economic profits or rents, adjusts reported earnings by deducting a charge for equity capital. Note that the market discount rates estimated from the abnormal earnings and dividend growth approaches are labeled differently: \( k \) and \( k^* \). Also, the standard transversality conditions apply to both models: in the limit as \( t \) approaches infinity, the present value of future price, \( p_t \) (difference between price and book value, \( p_t - bv_t \)) must tend to zero in equation (2) (in equation (4)).

Financial economists have expressed concerns about accounting earnings deviating from “true” earnings (and book values of equity deviating from market values), in the sense that accounting numbers are noisy and easily manipulated. However, the equivalence between equations (2) and (4) is not impaired by differences between accounting and economic numbers, nor is it affected by the latitude available within accounting rules to report different
accounting numbers. As long as forecasted earnings satisfy the clean surplus relation in equation (3) in terms of expectations, equation (4) is simply an algebraic restatement of equation (2), subject to the respective transversality conditions mentioned above.

Since the I/B/E/S database we use does not provide analysts’ earnings forecasts beyond year +5, we assume that abnormal earnings grow at a constant rate \( g_{ae} \) after year +5, to incorporate dates past that horizon. Equation (4) is thus adapted as follows.

\[
p_0 = b v_0 + \frac{a e_1}{(1 + k)} + \frac{a e_2}{(1 + k)^2} + \frac{a e_3}{(1 + k)^3} + \frac{a e_4}{(1 + k)^4} + \frac{a e_5}{(1 + k)^5} \left[ \frac{a e_5(1 + g_{ae})}{(k - g_{ae})(1 + k)^5} \right].
\]

The last, bracketed term is a terminal value that captures the present value of abnormal earnings after year +5. The terms before are derived from accounting statements \( b v_0 \) and analyst forecasts \( e_1 \) to \( e_5 \). Note that there are three separate growth rates in this paper and the different growth rates refer to different streams and periods and arise from different sources. The rate \( g \) refers to dividend growth in perpetuity and is assumed by the researcher; \( g_5 \) refers to growth in accounting earnings over the first five years and is provided by financial analysts; and \( g_{ae} \) refers to abnormal earnings growth past year +5 and is assumed by the researcher.

Whereas expected rates of return are typically viewed as being stochastic (Samuelson (1965)), \( k^* \) and \( k \) in equations (1) and (5) are nonstochastic discount rates. Barring a few recent exceptions (e.g., Ang and Liu (2000) and Vuolteenaho (1999)), the literature has assumed that expected rates of return can be approximated by discount rates. We make that assumption too. While equation (1) is designed to only reflect a flat \( k^* \), equation (2) can be restated to incorporate predictable variation over time in discount rates. Similarly, equation (5) can be restated to incorporate nonflat discount rates, as shown in Claus and Thomas (1999a). We consider the case when the equity premium is assumed to remain flat but discount rates vary over future periods based on the term-structure of risk-free rates. This restated version of equation (5) is

\[
p_0 = b v_0 + \sum_{t=1}^{\infty} \frac{a e_t}{\prod_{s=1}^{t} (1 + r_{fs} + rp)} ,
\]

where

\( r_{fs} = \) forward one-year risk-free rate for year \( s \),
\( rp = \) equity risk premium, assumed constant over all future years,
\( ae_t = \) expected abnormal earnings for year \( t \), equals \( e_t - bv_{t-1}(r_{fs} + rp) \) for years +1 through +5, and equals \( a e_5(1 + g_{ae})^{t-5} \), from year +6 on.
While the abnormal earnings stream in equation (4) is equivalent to the corresponding dividend stream in equation (2), the abnormal earnings relation in equation (5) (and equation (5a)) offers the following advantages over the dividend growth model in equation (1). First, a substantial fraction of the “value profile” for the abnormal earnings model in equation (5) is fixed by numbers that are currently available and do not need to be assumed by the researcher (current book value and abnormal earnings for years +1 through +5). Value profile is a representation of the fraction of total value captured by each future year’s flows. In contrast, the entire value profile for the dividend growth model is affected by the assumed growth rate, g. Since the fraction of value determined by assumed growth rates is lower for the abnormal earnings approach, those risk premium estimates are more reliable.

Second, in contrast to the potential for disagreement about a reasonable range for g, the rate at which rents can grow in perpetuity after year +5, \( g_{ae} \), is less abstract and easier to gauge using economic intuition. For example, to obtain equity premia around 8 percent, rents at the market level would have to grow forever at about 15 percent, on average. It is unlikely that aggregate rents to U.S. equity holders would grow at such high rates in perpetuity because of factors such as antitrust actions, global competition, and pressure from other stakeholders. The historical evidence (e.g., Myers (1999)) is also at odds with such high growth rates in abnormal earnings.

Third, future streams for a number of value-relevant indicators, such as price-to-book ratios \( P/B \), price-to-earnings ratios \( P/E \), and accounting return on equity \( roe \), can also be projected under the abnormal earnings approach. This allows one to paint a more complete picture of the future for different assumed growth rates. Analysis of the levels of future \( P/B \) and profitability (excess of \( roe \) over \( k \)) implied by growth rates required to obtain equity premium estimates around eight percent are also inconsistent with past experience.

II. Data and Methodology

I/B/E/S provides the consensus of all available individual forecasts as of the middle (the Thursday following the second Friday) of each month. Forecasts and prices should be gathered soon after the prior year-end, as soon as equity book values \( b_{v0} \) are available. Rather than collect forecasts at different points in the year, depending on the fiscal year-end of each firm, we opted to collect data as of the same month each year for all firms to ensure that the risk-free rate is the same across each annual sample. Since most firms have December year-ends, and book values of equity can be obtained from the balance sheets that are required to be filed with the SEC within 90 days of the fiscal year-end, we collect forecasts as of April each year.\(^8\) For

\(^8\) For the few firm-years not filing within this 90-day deadline, the book value of equity can be inferred by the market by adding (subtracting) fourth quarter earnings (dividends) from the third quarter book value of equity.
firms with fiscal year-ends other than December, this procedure creates a slight upward bias in estimated equity premium, since the stock prices used (as of April) are on average higher than those near the prior year's fiscal year-end, when $bv_0$ was released. In addition to earnings forecasts, I/B/E/S also provides data for actual earnings per share, dividends per share, share prices, and the number of outstanding shares. Equity book values are collected from COMPUSTAT's Industrial Annual, Research, and Full Coverage Annual Files, for years up to and including 1997.

The sample includes firms with I/B/E/S earnings forecasts for years $+1$ and $+2$ ($e_1$ and $e_2$) and a five-year growth forecast ($g_5$) as well as share prices and shares outstanding as of the I/B/E/S cut off date each April. We also require nonmissing data for the prior year's book value, earnings, and dividends. Explicit forecasts for years $+3$, $+4$, and $+5$ are often unavailable, and are generated by projecting the growth rate $g_5$ on the prior year's earnings forecast: $e_t = e_{t-1}(1 + g_5)$.9

Earlier years in the I/B/E/S database, before 1985, were dropped because they provided too few firms with complete data to represent the overall market. From 1985 on, the number of firms with available data increases substantially. As shown in column 1 of Table I, the number of sample firms increases from 1,559 in 1985 to 3,673 in 1998. Comparison with the total number of firms and market capitalization of all firms on NYSE, AMEX, and Nasdaq each April indicates that, although our sample represents only about 30 percent of all such firms, it represents 90 percent or more of the total market capitalization. Overall, we believe our sample is fairly representative of the value-weighted market, and refer to it as “the market” hereafter.

Firm-level data are aggregated each year to generate market-level earnings, book values, dividends, and capitalization. Actual data for year 0 (the full fiscal year preceding each April when forecasts were collected) is provided in columns 2 through 6 of Table I. Forecasted and projected earnings for years $+1$ through $+5$ are reported in columns 7 through 11.

Table I reveals an interesting finding relating to dividend payouts: the ratio of market dividends to earnings is around 50 percent in most years (with a noticeable decline toward the end of the sample period).10 We use this 50 percent payout ratio to project future dividends from earnings fore-

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9 If any of the explicit earnings forecasts for years $+2$, $+3$, $+4$, or $+5$ were negative, they were not used to project earnings for subsequent years. For about five percent of our sample, explicit earnings forecasts are available for all five years and do not need to be inferred using $g_5$. That subsample was investigated to confirm that projections based on five-year growth rates are unbiased proxies for the explicit forecasts for those years.

10 Although this statistic is well known to macroeconomists, it is higher than average firm-level dividend payouts. Note, however, that aggregate earnings include many loss firms, especially in the early 1990s, when earnings were depressed because of write-offs and accounting changes. This results in a higher aggregate dividend payout than the average firm-level payout ratio, which is computed over profitable firms only (the payout ratio is meaningless for loss firms). Also, since the aggregate payout ratio is a value-weighted average dividend payout, it is more representative of large firms, which tend to have higher dividend payouts than small firms.
Table I


The market consists of firms on the I/B/E/S Summary files with forecasts for years +1, +2, and a five-year earnings growth estimate ($g_e$) as of April each year, and actual earnings per share, dividends per share, number of shares outstanding and share prices as of the end of the prior fiscal year (year 0). Book values of equity for year 0 are obtained from COMPUSTAT. When missing on the I/B/E/S files, forecasted earnings per share for years +3, +4, and +5 are determined by applying $g_e$, the forecasted five-year growth rate, to year +2 forecasted earnings. All per share numbers are multiplied by the number of shares outstanding to get amounts at the firm level, and these are added across firms to get amounts at the market level each year. All amounts, except for dividend payout, are in millions of dollars.

<table>
<thead>
<tr>
<th>Forecast as of April</th>
<th>Number of Firms</th>
<th>Actual Values for Year 0</th>
<th>Forecast Earnings for Years +1 to +5</th>
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<tbody>
<tr>
<td></td>
<td></td>
<td>Earnings</td>
<td>Dividends</td>
</tr>
<tr>
<td>1985</td>
<td>1,559</td>
<td>154,858</td>
<td>71,134</td>
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<td>1986</td>
<td>1,613</td>
<td>155,201</td>
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<td>1987</td>
<td>1,774</td>
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<td>1988</td>
<td>1,735</td>
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<td>1991</td>
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<td>1998</td>
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</table>
casts, as well as to project future book values (using equation (3)). The validity of this assumption is not critical; however, varying the payout ratio between 25 and 75 percent has little impact on the estimated discount rate (results available upon request).

Both short- and long-term risk-free rates have been used in studies that estimate discount rates from flows that extend over many future periods. While one-month or one-year rates are appropriate when inferring the equity premium from historic returns (observed return less risk-free yield for that period), for studies based on forecasted flows, the maturity of risk-free rates used should match that of the future flows (Ibbotson Associates (1999)). Although we allow for expected variation in risk-free rates when estimating the risk premium, using equation (5a), we find almost identical results using a constant risk-free rate in equation (5) equal to the long-term rate. In essence, the shape of the yield curves over our sample period is such that the forward rates settle rather quickly at the long-term rate, and the impact of discounting flows from earlier years in the profile at rates lower than the long-term rate is negligible. For the sensitivity analyses, we find it convenient to use the constant rate structure of equation (5), rather than the varying rate structure of equation (5a). We selected the 10-year risk-free rate for the constant risk-free rate because it is the longest maturity for which data could be obtained for all country-years in our sample. To allow comparisons with other studies that use 30-year risk-free rates, we note that the mean 30-year risk-free rate in April for each year of our U.S. sample period is 31 basis points higher than the mean 10-year risk-free rate we use.

For years beyond year +5, abnormal earnings are assumed to grow at the expected inflation rate, \( g_{ae} \). As explained in the Appendix, the expected nominal inflation rate is higher than values of \( g_{ae} \) assumed in the literature, and is an upper bound for expected growth in abnormal earnings. We derive the expected inflation rate from the risk-free rate, based on the assumption that the real risk-free rate is approximately three percent.\(^{11}\) Since we recognize that this assumption is only an educated guess, we consider in Section V.D other values of \( g_{ae} \) also. Fortunately, our estimated risk premium is relatively robust to variation in the assumed growth rate, \( g_{ae} \), since a lower proportion of current market value is affected by \( g_{ae} \) in equations (5) and (5a), relative to the impact of \( g \) in equation (1).

### III. Results

Since \( k \) appears in both the numerators (\( ae \), is a function of \( k \)) and denominators of the terms on the right-hand side of equation (5), the resulting

\(^{11}\) The observed yields on recently issued inflation-indexed government bonds support this assumption. Although estimates of the real risk-free rate vary through time, and have historically been lower than three percent, more recently, the excess of the long-term risk-free rate over inflation forecasts has risen to three or four percent (e.g., Blanchard (1993), and discussion by Siegel).
equation is a polynomial in $k$ with many possible roots. Empirically, however, only one root is real and positive (see Botosan (1997)). We search manually for the value of $k$ that satisfies the relation each year, with the first iteration being close to the risk-free rate. The equity risk premium estimate ($rp$) that satisfies the valuation relation in equation (5a) is also estimated iteratively.

Table II provides the results of estimating $rp$, $k$, and $k^*$. The annual estimates for $rp$ (in column 13) lie generally between three and four percent and are much lower than the historic Ibbotson estimate. Also, there is little variation over time: each annual estimate is remarkably close to the mean value of 3.39 percent. The annual estimates for $k$ (in column 9) vary between a high of 14.38 percent in 1985 and a low of 8.15 percent in 1998. The corresponding risk-free rates (10-year Government T-bond yields) reported in column 8 vary with the estimated $k$s, between 11.43 percent in 1985 and 5.64 percent in 1998. As a result, the estimated equity premia (in column 11), equal to $k$ less $r_f$, exhibit little variation around the time-series mean of 3.40 percent.

While the equation (5a) equity premium estimates ($rp$) derived from non-flat risk-free rates are in concept more accurate than those derived by subtracting 10-year risk-free rates from the flat $k$ estimated from equation (5), the numbers reported in column 11 are very similar to those reported in column 13. We only consider the equation (5) estimates hereafter because (a) the magnitudes of the discount rates and their relation to risk-free rates are more transparent for the risk premium estimates based on constant risk-free rates, and (b) forward one-year rates for different maturities are not available for the other five markets.

To understand better the relative magnitudes of the terms in equation (5), we report in the first seven columns of Table II the fraction of market values represented by each term. The fraction represented by book value (column 1) has generally declined over our sample period, from 68.2 percent in 1985 to 26.4 percent in 1998. To compensate, the fraction represented by terminal value (column 7) has increased from 26.6 percent in 1985 to 60 percent in 1998. The fraction represented by abnormal earnings for years +1 to +5 has also increased.

Column 10 of Table II contains our estimates for $k^*$, the market discount rate based on the dividend growth model described by equation (1), when dividends are assumed to grow in perpetuity at the five-year growth in earnings forecast ($g_5$). Since $g_5$ is not available at the aggregate level, we use the forecast growth in aggregate earnings from year +4 to +5 (see column 16 of Table V) to identify $g_5$ at the market level. To maintain consistency with prior research using the dividend growth model, we estimate $d_1$ by applying the earnings growth forecast for year 1 on prior year dividends ($d_1 = d_0 \times e_1/e_0$). Our estimates for $k^*$ are almost identical to those reported by Moyer and Patel (1997).12 Note that these estimates of $k^*$ are much larger than the

---

12 Similar results are expected because the underlying data is taken from the same source, with minor differences in samples and procedures; for example, they use the S&P 500 index whereas we use all firms with available data.
Table II

Implied Expected Rate of Return on the Market ($k$ and $k^*$) and Equity Risk Premium ($rp$ and $k - r_f$) for U.S. Stocks (1985 to 1998)

The market is an aggregate of firms on the I/B/E/S Summary files with forecasts for years +1, +2, and a five-year earnings growth estimate ($g_5$) as of April each year, and actual earnings, dividends, number of shares outstanding and prices as of the end of the prior full fiscal year (year 0). Book values of equity for year 0 ($b_{v0}$) are obtained from COMPUSTAT. When missing, forecasted earnings for years +3, +4, and +5 are determined by applying $g_5$, the forecasted five-year growth rate, to year +2 forecasted earnings. The implied discount rate that satisfies the valuation relation in equation (5) below is $k$. Abnormal earnings ($ae_t$) equal reported earnings less a charge for the cost of equity (= beginning book value of equity * $k$). Assuming that 50 percent of earnings are retained allows the estimation of future book values from current book values and forecast earnings. The terminal value represents all abnormal earnings beyond year +5. Those abnormal earnings are assumed to grow at a constant rate, $gae$, which is assumed to equal the expected inflation rate, and is set equal to the current 10-year risk-free rate less 3 percent. The expected rate of return on the market is also estimated using equation (1), and is labeled $k^*$. Equation (1) is derived from the dividend growth model, and dividend growth in perpetuity, $g$, is assumed to equal the five-year earnings growth rate, $g_5$. Subtracting $r_f$ from the discount rates $k$ and $k^*$ generates equity premium estimates. The equity premium ($rp$) is also estimated using equation (5a), which is based on the same information used in equation (5), except that the constant discount rate $k$ is replaced by forward one-year risk-free rates at different maturities ($r_{fs}$) plus a constant risk premium ($rp$). All amounts, except for rates of return, are in millions of dollars.

\[
k^* = \frac{d_1}{p_0} + g
\]  

\[
p_0 = b_{v0} + \frac{ae_1}{(1 + k)} + \frac{ae_2}{(1 + k)^2} + \frac{ae_3}{(1 + k)^3} + \frac{ae_4}{(1 + k)^4} + \frac{ae_5}{(1 + k)^5} + \left[ \frac{ae_5(1 + g_{mr})}{(k - g_{mr})(1 + k)^5} \right]
\]  

\[
p_0 = b_{v0} + \sum_{t=1}^{\infty} \frac{ae_t}{\prod_{s=1}^{t} (1 + r_{fs} + rp)}
\]
<table>
<thead>
<tr>
<th>Forecast as of April</th>
<th>Book Value as Percent of Market Value</th>
<th>Percent of Market Value Represented by Present Value of Terminal Value</th>
<th>10-year $r_f$ from (5)</th>
<th>$k$ from (1)</th>
<th>$k^+$ from (1)</th>
<th>$k - r_f$</th>
<th>$k^+ - r_f$ from (5a)</th>
<th>$r_p$</th>
</tr>
</thead>
<tbody>
<tr>
<td>1985</td>
<td>68.2%</td>
<td>0.5% 0.9% 1.1% 1.3% 1.5% 26.6% 11.43% 14.38% 16.14% 2.95% 4.71% 2.88%</td>
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<tr>
<td>1986</td>
<td>53.2%</td>
<td>1.6% 2.0% 2.1% 2.3% 2.4% 36.3% 7.30% 11.28% 14.90% 3.98% 7.60% 4.03%</td>
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<tr>
<td>1987</td>
<td>50.1%</td>
<td>1.3% 1.9% 2.1% 2.2% 2.3% 40.0% 8.02% 11.12% 15.08% 3.10% 7.06% 3.25%</td>
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<tr>
<td>1988</td>
<td>54.7%</td>
<td>1.7% 1.8% 1.9% 2.0% 2.2% 35.7% 8.72% 12.15% 15.52% 3.43% 6.80% 3.58%</td>
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<tr>
<td>1989</td>
<td>53.9%</td>
<td>2.0% 2.0% 2.0% 2.1% 2.2% 35.7% 9.18% 12.75% 14.85% 3.57% 5.67% 3.54%</td>
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<tr>
<td>1990</td>
<td>52.0%</td>
<td>1.6% 2.0% 2.1% 2.2% 2.3% 37.8% 8.79% 12.33% 15.41% 3.54% 6.62% 3.56%</td>
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<tr>
<td>1991</td>
<td>48.5%</td>
<td>1.1% 1.9% 2.0% 2.2% 2.4% 41.8% 8.04% 11.05% 15.16% 3.01% 7.12% 2.96%</td>
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<tr>
<td>1992</td>
<td>47.8%</td>
<td>1.1% 1.9% 2.1% 2.3% 2.5% 42.4% 7.48% 10.57% 15.55% 3.09% 8.07% 3.06%</td>
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<tr>
<td>1993</td>
<td>43.5%</td>
<td>1.7% 2.3% 2.5% 2.7% 2.9% 44.4% 5.97% 9.62% 15.12% 3.65% 9.15% 3.76%</td>
<td></td>
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<td>1994</td>
<td>41.1%</td>
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<tr>
<td>1995</td>
<td>42.5%</td>
<td>2.1% 2.6% 2.7% 2.8% 3.0% 44.3% 7.06% 11.03% 14.96% 3.97% 7.90% 4.02%</td>
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<tr>
<td>1996</td>
<td>38.8%</td>
<td>2.2% 2.5% 2.6% 2.8% 3.0% 48.2% 6.51% 9.96% 14.96% 3.45% 8.45% 3.50%</td>
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<tr>
<td>1997</td>
<td>36.1%</td>
<td>2.2% 2.5% 2.6% 2.8% 3.0% 50.8% 6.89% 10.12% 13.88% 3.23% 9.99% 3.25%</td>
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<tr>
<td>1998</td>
<td>26.4%</td>
<td>2.1% 2.5% 2.7% 3.0% 3.2% 60.0% 5.64% 8.15% 13.21% 2.51% 7.57% 2.53%</td>
<td></td>
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<td></td>
</tr>
<tr>
<td>Mean</td>
<td></td>
<td></td>
<td>7.64% 11.04% 14.98% 3.40% 7.34% 3.39%</td>
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</tbody>
</table>
corresponding values of \( k \), and the implied equity premium estimates reported in column 12 \((k^* - r_f)\) are about twice those in column 11 \((k - r_f)\). The mean equity premium of 7.34 percent in column 12 of Table II is approximately the same as the Ibbotson estimate. Note also the larger variation in column 12, around this mean, relative to the variation in columns 11 and 13.

The results in Table II can be used to illustrate two primary advantages of the abnormal earnings model over the dividend growth model. First, the abnormal earnings approach uses more available “hard” data (current book value and forecast abnormal earnings for years +1 to +5) to reduce the emphasis on “softer” growth assumptions \((g_{ae})\) used to build terminal values. Figure 1 contains a value profile for the terms in equation (5), using data for 1991. This year was selected because it represents a “median” profile: the terminal value is a smaller (larger) fraction of total value for years before (after) 1991. Recall from Table II that our estimate for \( k \) in 1991 is 11.05 percent. The terminal value is based on abnormal earnings growing at an anticipated inflation rate of 5.04 percent \((g_{ae} \text{ is three percent less than the risk-free rate of 8.04 percent})\). The value profile for the abnormal earn-
nings model, represented by the solid columns in Figure 1, shows that approximately 50 percent of the total value is captured by current book value, 10 percent is spread over the abnormal earnings for the next five years, and about 40 percent remains in the terminal value. This last term is the only one affected by our growth assumption. In contrast, for the dividend growth model in equation (1), the dividend growth rate \( g \), which is assumed to equal the five-year analyst forecast for earnings growth \( g_5 = 12.12 \) percent, is the primary determinant of the estimated \( k^* \) (= 15.16 percent).

To offer a different perspective on why growth assumptions are more influential for projected dividends, relative to abnormal earnings, we converted the abnormal earnings profile in Figure 1 to an isomorphic value profile for dividends, represented by the hollow columns in Figure 1. (Note that these dividends refer to the flows underlying \( k \), from the abnormal earnings model, and are different from the flows underlying \( k^* \), the dividend growth model estimate.) The year +5 terminal value for the dividend profile in Figure 1 corresponds to a dividend growth in perpetuity of 6.8 percent.\(^{13} \) Even though the abnormal earnings and dividend profiles in Figure 1 correspond to the same underlying projections, the terminal value for the dividend profile represents almost 85 percent of total value. As a result, assumed dividend growth rates have a larger impact on estimated discount rates, relative to abnormal earnings growth rate assumptions. For example, doubling the assumed value of \( g_{ae} \) to 10 percent increases the estimated discount rate by only about two percentage points. In contrast, increasing the dividend growth assumption by one percentage point raises the estimated discount rate by almost the same amount.\(^{14} \)

The second major benefit of the abnormal earnings approach is that we can narrow the range of reasonable growth assumptions \( g_{ae} \), relative to the assumed growth rate for dividends \( g \). Since \( g \) is a hypothetical rate, it is not easy to determine whether 12.12 percent (the value of \( g \) underlying our 1991 estimate for \( k^* \)) is more or less reasonable than the 6.8 percent dividend growth in perpetuity (after year +5) implied by our abnormal earnings model projections. Fortunately, restating implied dividend growth rates in terms of terminal growth in abnormal earnings makes it easier to see why some dividend growth assumptions are unreasonable. The assumption that dividends grow at 12.12 percent implies that abnormal earnings past year +5 would need to grow in perpetuity at about 15 percent per year in equa-

\(^{13} \) This dividend growth rate is obtained by using equation (1) on projected market value in year +5, rather than current market values \( p_0 \) and the dividend in year six is the dividend in year +5 (= 50 percent of the earnings forecast for year +5) times the unknown growth rate. That is, solve for \( g \) in the relation \( p_5 = d_5 (1+g)/(k-g) \).

\(^{14} \) Note that in equation (1), changes in \( g \) increase \( k^* \) by exactly the same amount. For the dividend value profile in Figure 1, however, dividends for years +1 to +5 have been fixed by forecasted earnings and dividend payout assumptions. Therefore, increases in the dividend growth rate underlying the terminal value increase the estimated discount rate by a slightly smaller amount.
tion (5). This abnormal earnings growth rate corresponds to a real growth in rents of 10 percent (assumed long-term inflation rate is 5.04 percent), which is clearly an unreasonably optimistic assumption.

In sum, our estimates of the equity risk premium using the abnormal earnings approach are considerably lower than the Ibbotson rate, even though we believe the analyst forecasts we use, as well as the terminal growth assumptions we make, are optimistic. Adjusting for such optimism would lower our estimates further. While our estimates from the dividend growth approach are much closer to the Ibbotson rate, we believe they are biased upward because the assumed growth rate \( g = g_5 \) is too high an estimate for dividend growth in perpetuity. The estimates from the abnormal earnings approach are more reliable because we use more available information to reduce the importance of assumed growth rates, and we are better able to reject growth rates as being infeasible by projecting rents rather than dividends. Additional benefits of using the abnormal earnings approach are illustrated in Section V.

### IV. Equity Premium Estimates from Other Markets

Other equity markets offer a convenient opportunity to validate our domestic results. As long as the different markets are integrated with the United States and are of similar risk, those markets’ estimates should proxy for the equity premium in the United States. We replicated the U.S. analysis on five other important equity markets with sufficient data to generate reasonably representative samples of those markets. Only a summary of our results is provided here; details of those analyses are in Claus and Thomas (1999b). The six markets exhibit considerable diversity in performance and underlying fundamentals over our sample period. This across-market variation increases the likelihood that the estimates we obtain from each market offer independent evidence.

As with the U.S. data, earnings forecasts, actual earnings per share, dividends per share, share prices, and the number of outstanding shares are obtained from I/B/E/S. Book values of equity as of the end of year 0 are collected from COMPSTAT and Global Vantage for Canada and from Datastream for the remaining four countries. Unlike I/B/E/S and COMPSTAT, Datastream drops firms that are no longer active. While such deletions are less frequent outside the United States, only surviving firms are included in our sample. Fortunately, no bias is created in this study since we equate market valuations with contemporaneous forecasts, and do not track performance.\(^{15}\) Therefore, even if the surviving firms (included in our sample) performed systematically better or worse than firms that were dropped, our equity premium estimates are unbiased as long as market prices and earnings forecasts in each year are efficient and incorporate the same information.

\(^{15}\) Note that there is no “backfilling” in our sample, where prior years’ data for successful firms are entered subsequently.
All data are denominated in local currency. Currency risk is not an issue here, since it is present in the required rates of returns for both equities and government bonds. Thus the difference between the two rates should be comparable across countries.

We find that analysts’ forecasts in these five markets exhibit an optimism bias, similar to that observed in the United States. We considered other potential sources of measurement error in the forecasts, but are confident that any biases created by these errors are unlikely to alter our equity premium estimates much. For example, in Germany, earnings could be computed in as many as four different ways: GAAP per International Accounting Standards, German GAAP, DVFA, and U.S. GAAP.\textsuperscript{16} I/B/E/S employees indicated that they have been more successful at achieving consistency in recent years (all forecasts are on a DVFA basis), but they are not as certain about earlier years in their database. While differences in basis between forecast and actual items would affect analyst bias, they do not affect our estimates of market discount rates. Differences in basis across analysts contaminate the consensus numbers used, but the estimated market discount rates are relatively insensitive to changes in the near-term forecasts used.

To select the month of analysis for each country, we followed the same logic as that for the U.S. analysis. December was the most popular fiscal year-end for all countries except for Japan, where it was March. We then identified the period after the fiscal year-end by which annual earnings are required to be disclosed. This period differs across countries (see Table 1 in Alford et al. (1993)): it is three months for Japan and the United States, four months for France, six months for Canada and the United Kingdom, and eight months for Germany. We selected the month following the reporting deadline as the “sure to be disclosed” month to collect forecasts for any given year.

To include a country-year in our sample, we required that the total market value of all firms in our sample exceed 35 percent of the market value of “primary stock holdings” for that country, as defined by Datastream. Although we used a low hurdle to ensure that our sample contained contiguous years for all countries, a substantially greater proportion of the Datastream Market Index than our minimum hurdle is represented for most country-years.

The equity-premium estimates using the abnormal earnings and dividend growth approaches as well as the prevailing risk-free rates for different country-year combinations with sufficient data are reported in Table III. The number of years with sufficient firms to represent the overall market was highest for Canada (all 14 years between 1985 and 1998), and lowest for Japan (8 years). As with the U.S. sample, we use a 50 percent aggregate

\textsuperscript{16} The German financial analyst society, Deutsche Vereinigung für Finanzanalyse (DVFA), has developed a system used by analysts (and often by firms) to adjust reported earnings data to provide a measure that is closer to permanent or core earnings. The adjustment process uses both reported financial information as well as firms’ internal records. GAAP refers to Generally Accepted Accounting Principles or the accounting rules under which financial statements are prepared in different domiciles.
Table III
Implied Equity Premium Using Abnormal Earnings and Dividend Growth Approaches
\((k - r_f \text{ and } k^* - r_f)\) for International Stocks (1985 to 1998)

The market is an aggregate of firms on the I/B/E/S Summary files with forecasts for years +1, +2, and a five-year earnings growth estimate \((g_5)\) as of April each year, and actual earnings, dividends, number of shares outstanding, and prices as of the end of the prior full fiscal year (year 0). Book values of equity for year 0 \((b_{v0})\) are obtained from COMPUSTAT, Global Vantage, and Datastream. Forecasted earnings for years +3, +4, and +5 are determined by applying \(g_5\), the forecasted 5-year growth rate, to year +2 forecasted earnings. All amounts are measured in local currencies. \(r_f\) is the 10-year government bond yield. The implied discount rate that satisfies the valuation relation in equation (5) below is \(k\). Abnormal earnings \((aei)\) equal reported earnings less a charge for the cost of equity \((= \text{beginning book value of equity } \times k)\). Assuming that 50% of earnings are retained allows the estimation of future book values from current book values and forecast earnings. The terminal value represents all abnormal earnings beyond year +5. Those abnormal earnings are assumed to grow at a constant rate, \(g_{ae}\), which is assumed to equal the expected inflation rate, and is set equal to \(r_f\) less 3 percent. The expected rate of return on the market is also estimated using equation (1), and is labeled \(k^*\). Equation (1) is derived from the dividend growth model, and dividend growth in perpetuity, \(g\), is assumed to equal the five-year earnings growth rate, \(g_5\).

\[
p_0 = b_{v0} + \frac{ae_1}{(1 + k)} + \frac{ae_2}{(1 + k)^2} + \frac{ae_3}{(1 + k)^3} + \frac{ae_4}{(1 + k)^4} + \frac{ae_5}{(1 + k)^5} + \left[ \frac{ae_5 (1 + g_{ae})}{(k - g_{ae}) (1 + k)^5} \right]
\]

\[
k^* = \frac{d_1}{p_0} + g
\]
<table>
<thead>
<tr>
<th>Year</th>
<th>Canada</th>
<th>France</th>
<th>Germany</th>
<th>Japan</th>
<th>U.K.</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>$r_f$</td>
<td>$k - r_f$</td>
<td>$k^* - r_f$</td>
<td>$r_f$</td>
<td>$k - r_f$</td>
</tr>
<tr>
<td>1985</td>
<td>10.50%</td>
<td>4.41%</td>
<td>7.45%</td>
<td>1986</td>
<td>8.82%</td>
</tr>
<tr>
<td></td>
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</tr>
</tbody>
</table>
The abnormal earnings estimates generate projections that are consistent with experience, but the dividend growth estimates are biased upward and generate projections that are too optimistic because the five-year earnings growth forecast \( g_5 \) is too high an estimate for dividend growth in perpetuity. The values of \( g_5 \) suggest mean real dividend growth rates in perpetuity that range between 6.09 percent for Canada and 8.25 percent for Japan. These real rates exceed historic real earnings growth rates, and are at least twice as high as the real GDP growth rates forecast for these countries.

The results observed for Japan are unusual and invite speculation. While our results suggest that the equity premium in Japan increased during the sample period, from about \(-1\) percent in the early 1990s to 2 percent in the late 1990s, these results are also consistent with a stock market bubble that has gradually burst. That is, early in our sample period, prices were systematically higher than the fundamentals (represented by analysts’ forecasts) would suggest, and have gradually declined to a level that is supported by analysts’ forecasts. Note that our sample excludes the peak valuations in the late 1980s before the crash. Perhaps the implied equity premium in that period would be even more negative than the numbers we estimate for the early 1990s. Regardless of whether the poor performance of Japanese equities in the 1990s is due to correction of an earlier mispricing, it is useful to contrast the inferences from a historic approach with those from a forward-looking approach such as ours: the former would conclude that equity premia have fallen in Japan during the 1990s, whereas our approach suggests the opposite.

V. Sensitivity Analyses

This section summarizes our analysis of U.S. equity data designed to gauge the robustness of our conclusion that the equity premium is much lower..
than historic estimates. We begin by considering two relations for P/B and P/E ratios that allow us to check whether our projections under the dividend growth and abnormal earnings models are reasonable. Next, we document the extent of analyst optimism in our data. Finally, we consider the sensitivity of our risk premium estimates to the assumed abnormal earnings growth rate ($g_{ae}$).17

### A. P/B Ratios and the Level of Future Profitability

The first relation we examine is that between the P/B ratio and future levels of profitability (e.g., Penman 1999), where future profitability is the excess of the forecast market accounting rate of return ($\text{roet}$) over the required rate of return ($k$).

$$
\frac{p_0}{b_{v_0}} = 1 + \frac{\text{roe}_1 - k}{(1 + k)} + \frac{\text{roe}_2 - k}{(1 + k)^2} \left( \frac{b_{v_1}}{b_{v_0}} \right) + \frac{\text{roe}_3 - k}{(1 + k)^3} \left( \frac{b_{v_2}}{b_{v_0}} \right) + \ldots,
$$

where $\text{roe}_t = \frac{e_t}{b_{v_{t-1}}}$ is the accounting return on equity in year $t$.

This relation indicates that the P/B ratio is explained by expected future profitability ($\text{roe}_t - k$).18 Firms expected to earn an accounting rate of return on equity equal to the cost of capital should trade currently at book values ($p_0/b_{v_0} = 1$). Similarly, the P/B ratio expected in year $+5$ ($p_{5}/b_{v_5}$), which is determined by the assumed growth in abnormal earnings after year $+5$ ($g_{ae}$), should be related to profitability beyond year $+5$. To investigate the validity of our assumed growth rates, we examine the profiles of future P/B ratios and profitability levels to check if they are reasonable and related to each other as predicted by equation (6). Future book values are generated by adding projected earnings and subtracting projected dividends (assuming a 50 percent payout) to the prior year’s book value. Similarly, projected market values are obtained by growing the prior year’s market value at the discount rate ($k$) less projected dividends.

Table IV provides data on current and projected values of P/B ratios and profitability. Current market and book values are reported in columns 1 and 2, and projected market and book values in year $+5$ are reported in columns

---

17 We also examined Value Line data for the DOW 30 firms for two years: 1985 and 1995 (details in Claus and Thomas 1999a). Value Line provides both dividend forecasts over a four- or five-year horizon and a projected price. This price is, in effect, a terminal value estimate, which obviates the need to assume dividend growth in perpetuity. Unfortunately, those risk premium estimates appear to be unreliable: The estimated discount rate is 20 percent (8.5 percent) for 1985 (1995). These results are consistent with Value Line believing that the DOW 30 firms are undervalued (overvalued) in 1985 (1995); that is, current price does not equal the present value of forecast dividends and projected prices. This view is supported by their recommendations for the proportion to be invested in equity: it was 100 percent through the 1980s, and declined through the 1990s (it is currently at 40 percent).

18 The growth in book value terms in equation (6), $b_{v_t}/b_{v_0}$, which add a multiplicative effect, have been ignored in the discussion because of the built-in correlation with $\text{roe}_t - k$. Higher $\text{roe}_t$ results in higher $e_t$, which in turn causes higher growth in $b_{v_t}$ because dividend payouts are held constant at 50 percent for all years.
Table IV
Price-to-Book Ratios ($p_t/bv_t$), Forecast Accounting Return on Equity ($roe_t$) and Expected Rates of Return ($k$) for U.S. Stocks (1985 to 1998)

To examine the validity of assumptions underlying $k$, which is the implied discount rate that satisfies the valuation relation in equation (5), current price-to-book ratios are compared with estimated future returns on equity ($roe_t$) to examine fit with equation (6) below. The market is an aggregate of firms on the I/B/E/S Summary files with forecasts for years +1, +2, and a five-year earnings growth estimate ($g_5$) as of April each year, and actual earnings, dividends, number of shares outstanding, and prices as of the end of the prior full fiscal year (year 0). Book values of equity for year 0 ($bv_0$) are obtained from COMPUSTAT. When missing, forecasted earnings for years +3, +4, and +5 are determined by applying $g_5$ to year +2 forecasted earnings. Assuming that 50 percent of earnings are retained allows the estimation of future book values from current book values and forecast earnings. Return on equity ($roe_t$) equals forecast earnings scaled by beginning book value of equity ($bv_{t-1}$). Market and book value amounts are in millions of dollars.

\[
p_0 = bv_0 + \frac{ae_1}{(1 + k)} + \frac{ae_2}{(1 + k)^2} + \frac{ae_3}{(1 + k)^3} + \frac{ae_4}{(1 + k)^4} + \frac{ae_5}{(1 + k)^5} + \left[ \frac{ae_5(1 + g_5)}{(k - g_5)(1 + k)^5} \right]
\]  

\[
\frac{p_0}{bv_0} = 1 + \frac{roe_1 - k}{(1 + k)} + \frac{roe_2 - k}{(1 + k)^2} \left( \frac{bv_1}{bv_0} \right) + \ldots
\]
<table>
<thead>
<tr>
<th>Year 0 Equity Values</th>
<th>Year +5 Equity Values</th>
<th>Price/Book Ratio</th>
<th>Forecast Accounting Return on Equity</th>
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<tbody>
<tr>
<td></td>
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<td></td>
</tr>
<tr>
<td><strong>Forecasts as of April</strong></td>
<td><strong>Market Value</strong></td>
<td><strong>Book Value</strong></td>
<td><strong>Market Value</strong></td>
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<tr>
<td></td>
<td>$p_0$</td>
<td>$b_0$</td>
<td>$p_5$</td>
</tr>
<tr>
<td>1985</td>
<td>1,747,133</td>
<td>1,191,869</td>
<td>2,676,683</td>
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<tr>
<td>1986</td>
<td>2,284,245</td>
<td>1,214,454</td>
<td>3,197,490</td>
</tr>
<tr>
<td>1987</td>
<td>2,640,743</td>
<td>1,323,899</td>
<td>3,727,459</td>
</tr>
<tr>
<td>1988</td>
<td>2,615,857</td>
<td>1,430,672</td>
<td>3,779,033</td>
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<tr>
<td>1989</td>
<td>2,858,585</td>
<td>1,541,231</td>
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<tr>
<td>1990</td>
<td>3,143,879</td>
<td>1,636,069</td>
<td>4,589,685</td>
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<td>1991</td>
<td>3,660,296</td>
<td>1,775,199</td>
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</tr>
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<td>4,001,756</td>
<td>1,911,383</td>
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<td>4,918,359</td>
<td>2,140,668</td>
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<td>1994</td>
<td>5,282,046</td>
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<td>1995</td>
<td>6,289,760</td>
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<td>8,207,274</td>
<td>3,182,952</td>
<td>11,206,787</td>
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<td>1997</td>
<td>10,198,036</td>
<td>3,679,110</td>
<td>14,103,523</td>
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<tr>
<td>1998</td>
<td>12,908,495</td>
<td>3,412,303</td>
<td>16,838,377</td>
</tr>
<tr>
<td><strong>Mean</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>
3 and 4. These values are used to generate current and year +5 P/B ratios, reported in columns 5 and 6. Columns 7 through 12 contain the forecasted accounting rate of return on equity for years 1 to 6, which can be compared with the estimated market discount rate, $k$, reported in column 13, to obtain forecasted profitability.

The current P/B ratio has been greater than 1 in every year in the sample period, and has increased steadily over time, from 1.5 in 1985 to 3.8 in 1998. Consistent with equation (6), all forecasted roe values for years 1 through 6 in Table IV exceed the corresponding values of $k$. Increases in the P/B ratio over the sample period are mirrored by corresponding increases in forecast profitability ($\text{roe}_t - k$) in years +1 through +5 as well as forecast profitability in the posthorizon period (after year +5), as measured by the implied price-to-book ratio in year +5. Finally, the tendency for P/B ratios to revert gradually over the horizon toward one (indicated by the year +5 values in column 6 being smaller than the year 0 values in column 5) is consistent with intuition (e.g., Nissim and Penman (1999)).

We also extended our investigation to years beyond year +5 for the assumptions underlying the abnormal earnings estimates, and find that the pattern of projections for P/B and roe remain reasonable. In contrast, those projections for the assumptions underlying the dividend growth model estimates suggest that the underlying growth rates are unreasonably high. To provide an illustrative example of those results, we contrast in Figure 2 the patterns for future roe and P/B that are projected for the dividend growth and abnormal earnings approaches for 1991. The roe levels are marked off on the left scale, and P/B ratios are shown on the right scale. Recall that the market discount rates estimated for the abnormal earnings and dividend growth approaches are 11.05 percent ($k$) and 15.16 percent ($k^*$) and the corresponding terminal growth rates for abnormal earnings and dividends are 5.04 percent and 12.12 percent.

The projections for the abnormal earnings method (indicated by bold lines) continue to remain reasonable. The P/B ratio always exceeds one, but it trends down over time. Consistent with P/B exceeding one, the roe is always above the 11.05 percent cost of capital, and trends toward it after year +5. Note that the optimistic analyst forecasts cause roe projections to climb for years +1 through +5, but the subsequent decline in roe is because the profitability growth implied by $g_{ae}$ (our assumed growth in abnormal earnings past year +5) is lower than that implied by $g_5$.

The results for the dividend growth approach illustrate the benefits of using projected accounting ratios to validate assumed growth rates. The profitability (roe) is actually below the cost of equity of 15.16 percent ($k^*$), for the first three years, even though the P/B ratio is greater than one. Thereafter, the profitability keeps increasing, to a level above 20 percent by year +15. Both the high level of profitability and its increasing trend are not easily justified, especially when they are observed repeatedly for every year in our sample. Similarly, the increasing pattern for P/B, which is projected to increase from about two to about three by year +15, is hard to justify.
These projections are, however, consistent with an estimated discount rate that is too high. Since near-term analysts’ forecasts of profitability are below this discount rate, future levels of profitability have to be unreasonably high to compensate.

B. P/E Ratios and Forecast Growth in Profitability

The second relation we use to check the validity of our assumptions regarding $g_{ae}$ is the price–earnings ratio, described by equation (7) (see derivation in Claus and Thomas, 1999a). Price–earnings ratios are a function of the present value of future changes in abnormal earnings, multiplied by a capitalization factor ($= 1/k$).

$$
\frac{P_0}{e_t} = \frac{1}{k} \left[ 1 + \frac{\Delta ae_2}{e_t(1+k)} + \frac{\Delta ae_3}{e_t(1+k)^2} + \ldots \right],
$$

where $\Delta ae_t = ae_t - ae_{t-1}$ is the change in expected abnormal earnings over the prior year.
The price–earnings ratio on the left-hand side deviates slightly from the traditional representation in the sense that it is a “forward” price–earnings ratio, based on expected earnings for the upcoming year, rather than a “trailing” price–earnings ratio \( \frac{p_0}{e_0} \), which is based on earnings over the year just concluded. The relation between future earnings growth and forward price–earnings ratios is simpler than that for trailing price–earnings ratios.\(^{19}\) Therefore, we use only the forward price–earnings ratio here and refer to it simply as the P/E ratio.

The results reported in Table V describe P/E ratios and growth in abnormal earnings derived from analysts’ forecasts for the market. The first four columns provide market values and the corresponding upcoming expected earnings for year 0 and year +5. These numbers are used to generate the current and year +5 P/E ratios reported in columns 5 and 6, which can be compared to the values of \( 1/k \) reported in column 18.\(^{20}\) According to equation (7), absent growth in abnormal earnings, the P/E ratio should be equal to \( 1/k \), and the P/E ratio should be greater (less) than \( 1/k \) for positive (negative) expected growth in abnormal earnings. Forecast growth rates in abnormal earnings for years +2 through +6 are reported in columns 7 through 11. To maintain equivalence with the terms in equation (7), growth in abnormal earnings is scaled by earnings expected for year +1 \( (e_1) \) and then discounted.

To understand the relations among the numbers in the different columns, consider the row corresponding to 1991. The market P/E ratio of 15.1 is higher than the inverse of the discount rate \( (1/k = 9.0) \). That difference of 6.1 is represented by the sum of the present value of the abnormal earnings growth terms in future years, scaled by \( e_1 \) (this sum needs to be multiplied by \( 1/k \) as shown in equation (7)). These growth terms decline from 13 percent in year 2 to 2 percent in year 6, and continue to decline thereafter. By year +5, the market P/E is expected to fall (to 11.7), since some of the growth in abnormal earnings (represented by the amounts in columns 7 through 11) is expected to have already occurred by then. Turning to the other sample years, the P/E ratios in year 0 (column 5) have generally increased through the sample period, and so have the values of \( 1/k \). Consistent with P/E ratios exceeding \( 1/k \) in every year, abnormal earnings are forecast to exhibit positive growth for all cells in columns 7 to 11. Also, the P/E ratios in year +5 are forecast to decline, relative to the corresponding year 0 P/E values, because of the value represented by the amounts in columns 7 to 11.

\(^{19}\) Since the numerator of the P/E ratio is an ex-dividend price \( (p_0) \), the payment of a large dividend \( (d_0) \) would reduce \( p_0 \) without affecting trailing earnings \( (e_0) \), thereby destroying the relation between \( p_0 \) and \( e_0 \). This complication does not arise when expected earnings for the upcoming period \( (e_1) \) is used instead of \( e_0 \).

\(^{20}\) If the numbers in Table V appear to be not as high as the trailing P/E ratios commonly reported in the popular press, note that forward P/E ratios are generally smaller than trailing P/E ratios for the following reasons. First, next year’s earnings are greater than current earnings because of earnings growth. Second, current earnings contain one-time or transitory components that are on average negative, whereas forecast earnings focus on core or continuing earnings.
For purposes of comparison with other work, we also report in columns 12 through 17 of Table V the growth in forecast earnings (as opposed to growth in abnormal earnings) for years +1 through +6. Forecasted growth in earnings declines over the horizon, similar to the pattern exhibited by growth in abnormal earnings. Note the similarity in the pattern of earnings growth for all years in the sample period: the magnitudes of earnings growth estimates appear to settle at around 12 percent by year +5, before dropping sharply to values around 7 percent in the posthorizon period (year +6). Again, this decline occurs because the earnings growth implied by $g_{ae}$ (our assumed growth in abnormal earnings past year +5) is lower than $g_5$.

The results in Table V confirm the predictions derived from equation (7) as well as the intuitive links drawn in the literature. As with the results for P/B ratios, the trends for P/E ratios and growth in abnormal earnings exhibit no apparent discrepancies that might suggest that the assumptions underlying our abnormal earnings model are unreasonable.

C. Bias in Analyst Forecasts

We considered a variety of biases that may exist in the I/B/E/S forecasts, but found only the well-known optimism bias to be noteworthy (details provided in Claus and Thomas (1999a)).\textsuperscript{21} We compute the forecast error for each firm in our sample, representing the median consensus forecast as of April less actual earnings, for different forecast horizons (year +1, +2, \ldots +5) for each year between 1985 and 1997. Table VI contains the median forecast errors (across all firms in the sample for each year), scaled by share price. In general, forecasted earnings exceed actual earnings, and the extent of optimism increases with the horizon.\textsuperscript{22} There is, however, a gradual reduction in optimism toward the end of the sample period.

Since the forecast errors in Table VI are scaled by price, comparing the magnitudes of the median forecast errors with the inverse of the trailing P/E ratios (or E/P ratios) is similar to a comparison of forecast errors with earnings levels. While the trailing E/P ratios for our sample vary between 5 and 9 percent, the forecast errors in Table VI vary between values that are in the neighborhood of 0.5 percent for year +1 to around 3 percent in year +5. Comparing the magnitudes of year +5 forecast errors with the implied E/P ratios indicates that forecasted earnings exceed actual earnings by as

\textsuperscript{21} I/B/E/S removes one-time items (typically negative) from reported earnings. That is, the level of optimism would have been even higher if we had used reported numbers instead of actual earnings according to I/B/E/S.

\textsuperscript{22} In addition to increasing with forecast horizon, the optimism bias is greater for certain years where earnings were depressed temporarily. The higher than average dividend payouts observed in Table I for 1987 and 1992 indicate temporarily depressed earnings in those years, and the forecast errors are also higher than average for those years. For example, the two largest median year +2 forecast errors are 1.86 and 1.81 percent, and they correspond to two-year out forecasts made in 1985 and 1990.
Table V
Forward Price-to-Earnings Ratios ($p_t/e_{t+1}$) and Growth in Forecast Abnormal Earnings and Earnings for U.S. Stocks (1985 to 1998)

To examine the validity of assumptions underlying $k$, which is the implied discount rate that satisfies the valuation relation in equation (5), current and forecast forward price-to-earnings ratios are compared with growth in forecast abnormal earnings to examine fit with equation (7) below. The market is an aggregate of firms on the I/B/E/S Summary files with forecasts for years +1, +2, and a five-year earnings growth estimate ($g_5$) as of April each year, and actual earnings, dividends, number of shares outstanding, and prices as of the end of the prior full fiscal year (year 0). Book values of equity for year 0 ($b_{0t}$) are obtained from COMPUSTAT. Abnormal earnings ($ae_t$) equal reported earnings less a charge for the cost of equity ($-\text{beginning book value of equity} \times k$). Future market values are projected for each year by multiplying beginning market values by $(1 + k)$ and subtracting dividends. When missing, forecasted earnings for years +3, +4, and +5 are determined by applying $g_5$ to year +2 forecasted earnings. Assuming that 50 percent of earnings are retained allows the estimation of future book values from current book values and forecast earnings. Market equity values and earnings amounts are in millions of dollars.

\begin{equation}
p_0 = b_{0t} + \frac{ae_1}{(1 + k)} + \frac{ae_2}{(1 + k)^2} + \frac{ae_3}{(1 + k)^3} + \frac{ae_4}{(1 + k)^4} + \frac{ae_5}{(1 + k)^5} + \frac{ae_5(1 + g_5)}{(k - g_5)(1 + k)^5}
\end{equation}

\begin{equation}
\frac{p_0}{e_1} = \frac{1}{k} \left[ 1 + \frac{\Delta ae_2}{e_1(1 + k)} + \frac{\Delta ae_3}{e_1(1 + k)^2} + \ldots \right]
\end{equation}
<table>
<thead>
<tr>
<th>Year 0 Values</th>
<th>Year +5 Values</th>
<th>Forward P/E Ratio</th>
<th>PV of ae Growth ($\Delta e_t$), Scaled by $e_1$</th>
<th>Growth in Forecast Earnings</th>
<th>1/b from Eq. (5)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Market Value ($p_0$)</td>
<td>Earnings ($e_1$)</td>
<td>Market Value ($p_0$)</td>
<td>Earnings ($e_4$)</td>
<td>In Year 0</td>
</tr>
<tr>
<td></td>
<td>1985 1,747,133</td>
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<td>308,308</td>
<td>9.7</td>
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<td></td>
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<td></td>
<td>1997 10,196,036</td>
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<td></td>
<td>1998 12,908,495</td>
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<td>16,838,377</td>
<td>1,069,786</td>
<td>22.4</td>
</tr>
<tr>
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<td></td>
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<td></td>
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<td>14.6</td>
</tr>
</tbody>
</table>

The following table represents the median of all forecast errors scaled by share price for each year examined. The forecast error is calculated for each firm as of April each year, and equals the median consensus forecasted earnings per share minus the actual earnings per share, scaled by price. The year when the forecasts were made is listed in the first row, while the first column lists the horizon of that forecast. For each year and horizon combination, we report the median forecast error and the number of firms in the sample. To interpret the Table, consider the values of 0.78 percent and 1,680 reported for the +1/1985 combination, in the top left-hand corner of the table. This means that the median value of the difference between the forecasted and actual earnings for 1986 was 0.78 percent of price, and that sample consisted of 1,680 firms with available forecast errors. The results confirm that analyst forecasts are systematically positively biased and that this bias increases with the forecast horizon; however, the extent of any such bias has been declining steadily over time.

<table>
<thead>
<tr>
<th></th>
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<tbody>
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<td>Median</td>
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<td>0.37%</td>
<td>0.07%</td>
<td>0.44%</td>
<td>0.58%</td>
<td>0.39%</td>
<td>0.17%</td>
<td>0.15%</td>
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<td>Obs.</td>
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<td>0.79%</td>
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<td>—</td>
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<tr>
<td>Year +2</td>
<td>Median</td>
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<td>0.99%</td>
<td>1.44%</td>
<td>2.22%</td>
<td>2.78%</td>
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<td>0.54%</td>
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<td>2,159</td>
<td>2,396</td>
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<td>—</td>
<td>—</td>
<td>—</td>
</tr>
<tr>
<td>Year +3</td>
<td>Median</td>
<td>2.63%</td>
<td>2.04%</td>
<td>2.80%</td>
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<td>3.17%</td>
<td>2.83%</td>
<td>1.54%</td>
<td>0.91%</td>
<td>0.77%</td>
<td>0.60%</td>
<td>—</td>
<td>—</td>
<td>—</td>
</tr>
<tr>
<td>Obs.</td>
<td>1,285</td>
<td>1,344</td>
<td>1,492</td>
<td>1,474</td>
<td>1,586</td>
<td>1,696</td>
<td>1,724</td>
<td>1,825</td>
<td>2,024</td>
<td>2,132</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
</tr>
<tr>
<td>Year +4</td>
<td>Median</td>
<td>3.54%</td>
<td>3.44%</td>
<td>3.86%</td>
<td>3.59%</td>
<td>3.43%</td>
<td>2.91%</td>
<td>1.36%</td>
<td>0.94%</td>
<td>0.74%</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
</tr>
<tr>
<td>Obs.</td>
<td>1,201</td>
<td>1,260</td>
<td>1,411</td>
<td>1,432</td>
<td>1,528</td>
<td>1,621</td>
<td>1,618</td>
<td>1,704</td>
<td>1,815</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
</tr>
</tbody>
</table>
much as 50 percent at that horizon. These results suggest that our equity premium estimates are biased upward because we do not adjust for the considerable optimism in earnings forecasts for years +1 to +5. They also suggest that we are justified in dropping assumed growth rates for earnings past year +5 (column 17 versus column 16 in Table V).

D. Impact of Variation in the Assumed Growth Rate in Abnormal Earnings Beyond Year +5 ($g_{ae}$)

We begin by considering two alternative cases for $g_{ae}$: three percent less and three percent more than our base case, where $g_{ae}$ is assumed to equal the expected inflation rate. As mentioned in the Appendix, our base growth rate of $g_{ae} = r_f - 3\%$ is higher than any rate assumed in the prior abnormal earnings literature. Adding another three percent to the growth rate, which would require rents to grow at a three percent real rate in perpetuity, raises the level of optimism further. Dropping three percent from the base case, in the lower growth scenario, would be equivalent to assuming a very low nominal growth rate in abnormal earnings, and would be only slightly more optimistic than the assumptions in much of the prior abnormal earnings literature.

For the higher (lower) growth rate scenario, corresponding to $g_{ae} = r_f (g_{ae} = r_f - 6\%)$, the average risk premium over the 14-year sample period increases (decreases) to a mean of 4.66 (2.18), from a mean of 3.40 percent for the base case. Even for the high growth rate in abnormal earnings, the increase in the estimated risk premium is modest, and leaves it substantially below the traditional estimates of the risk premium. While increasing (decreasing) the growth rate increases (decreases) the terminal value, it also reduces (increases) the present value of that terminal value because of the higher (lower) discount rate it engenders.

We also considered a synthetic market portfolio each year constructed to have no expected future abnormal earnings, to avoid the need for an assumed abnormal earnings growth rate beyond year +5. As described in equation (6), portfolios with $P/B = 1$ should exhibit no abnormal earnings; that is, the $roe_t$ should on average equal $k$ for this synthetic market. The last term in equation (5), representing the terminal value of abnormal earnings beyond year +5, is set to zero and the estimates for $k$ obtained iteratively each year. The mean estimate for $k - r_f$ from this synthetic market is 2.20 percent, which is slightly lower than the mean risk premium of 3.40 percent in Table II. Note that a lower discount rate is not expected for the synthetic market, since it has a beta close to one each year and has a lower $P/B$ than the market. (Low $P/B$ firms are expected to generate higher returns (e.g., Gebhardt, Lee, and Swaminathan forthcoming).) The higher discount rates observed for the assumptions underlying our abnormal earnings model support our view that the analyst forecasts we use and our assumption that the terminal growth in abnormal earnings equals expected inflation ($g_{ae} = r_f - 3\%$) are both optimistic.
VI. Conclusion

Barring some notable exceptions (e.g., Siegel (1992 and 1998), Blanchard (1993), Malkiel (1996), and Cornell (1999)), academic financial economists generally accept that the equity premium is around eight percent, based on the performance of the U.S. market since 1926. We claim that these estimates are too high for the post-1985 period that we examine, and the equity premium is probably no more than three percent. Our claim is based on estimates of the equity premium obtained for the six largest equity markets, derived by subtracting the 10-year risk-free rate from the discount rate that equates current prices to forecasted future flows (derived from I/B/E/S earnings forecasts). Growth rates in perpetuity for dividends and abnormal earnings need to be much higher than is plausible to justify equity premium estimates of about eight percent. Not only are such growth rates substantially in excess of any reasonable forecasts of aggregate growth (e.g., GDP), the projected streams for various indicators, such as price-to-book and price-to-earnings ratios, are also internally contradictory and inconsistent with intuition and past experience.

We agree that the weight of the evidence provided by the historical performance of U.S. stock markets since 1926 is considerable. Yet there are reasons to believe that this performance exceeded expectations, because of potential declines in the equity premium, good luck, and survivor bias. While projecting dividends to grow at earnings growth rates forecast by analysts provides equity premium estimates as high as eight percent, we show that those growth forecasts exhibit substantial optimism bias and need to be adjusted downward. In addition to our results, theory-based work, historical evidence from other periods and other markets, and surveys of institutional investors all suggest that the equity premium is much lower than eight percent. Overall, we believe that an eight percent equity premium is not supported by an analysis that compares current market prices with reasonable expectations of future flows for the markets and years that we examine.

Appendix: Assumed Growth Rates in Perpetuity for Dividends ($g$) and Abnormal Earnings ($g_{ae}$)

While the conceptual definition of $g$ is clear—it is the dividend growth rate that can be sustained in perpetuity, given current capital and future earnings—determining this rate from fundamentals is not easy. To illustrate, take two firms that are similar in every way, except that they have announced different dividend policies in the current period, which results in a higher expected forward dividend yield ($d_1/p_0$) for one firm than the other, say 7 percent and 1 percent. What can be said about $g$ for the two firms?

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23 Assuming too high a rate would cause the capital to be depleted in some future period, and assuming too low a rate would cause the capital to grow “too fast.”
Examination of equation (1) indicates that $g$ for the low dividend yield firm must be 6 percent higher than $g$ for the higher dividend yield firm, assuming they both have the same discount rate ($k^*$). If $k^*$ equals 10 percent, for example, the value of $g$ for the two firms must be 3 percent and 9 percent. These two values of $g$ are substantially different from each other, even though the two firms are not.

In addition to being a hypothetical rate, $g$ need not be related to historic or forecasted near-term growth rates for earnings or dividends. Dividend payout ratios can change over time because of changes in the investment opportunity set available and the relative attractiveness of cash dividends versus stock buybacks. Since changes in dividend payout affect the dividend yield, which in turn affects $g$, historic growth rates may not be relevant for $g$. Also, if dividend policies are likely to change over time, $g$ need not be related to $g_5$ (the growth rate forecast for earnings over the next five years), a rate that is frequently used to proxy for $g$. Various scenarios can be constructed for the two firms in the example above to obtain similar historic and/or near-term forecast growth rates and yet have substantially different values for $g$.

Despite the difficulties noted above, both historic and forecast rates for aggregate dividends, earnings, and other macroeconomic measures (such as GDP) have been used as proxies for $g$. We note that these proxies create additional error. First, it is important to hold the unit of investment constant through the period where growth is measured. In particular, any growth created at the aggregate level by the issuance/retirement of equity since the beginning of the period should be ignored. Second, profits from all activities conducted outside the publicly traded corporate sector that are included in the macroeconomic measures should be deleted, and all overseas profits relating to this sector that are excluded from some macroeconomic measures should be included.

To control for the unit of investment problem, we use forecasted growth in per-share earnings rather than aggregate earnings, and to mitigate the problems associated with identifying $g$, we focus on growth in rents (abnormal earnings), $g_{ae}$, rather than dividends. To understand the benefits of switching to $g_{ae}$, it is important to describe some features of abnormal earnings. Expected abnormal earnings would equal zero if book values of equity reflected market values. If book values measure input costs fairly, but do not include the portion of market values that represent economic rents (not yet earned), abnormal earnings would reflect those rents. However, the magnitude of such rents at the aggregate market level is likely to be small, and any rents that emerge are likely to be dissipated over time for the usual reasons (antitrust actions, global competition, etc.). As a result, much of the

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24 That is, if market prices are efficient and book values are marked to market values each period, market (book) values are expected to adjust each period so that no future abnormal returns (abnormal earnings) are expected.
earlier literature using the abnormal earnings approach has assumed zero growth in abnormal earnings past the “horizon” date.\(^{25}\)

Returning to the two-firm example, shifting the focus from growth in dividends to growth in rents removes much of the confusion caused by transitory changes in dividend payouts and dividend yields: these factors should have no impact on growth in rents, since the level of and growth in rents are determined by economic factors such as monopoly power. That is, even though the two firms have different forecasted earnings and dividends, the forecasted abnormal earnings and growth in abnormal earnings should be identical.

We believe, however, that the popular assumption of zero growth in abnormal earnings may be too pessimistic because accounting statements are conservative and understate input costs: assets (liabilities) tend to be understated (overstated) on average. For example, many investments (such as research and development, advertising, and purchased intangibles) are written off too rapidly in many domiciles. As a result, abnormal earnings tend to be positive, even in the absence of economic rents. Growth in abnormal earnings under conservative accounting is best understood by examining the behavior of the excess of roe (the accounting rate of return on the book value of equity) over \(k\) (the discount rate). Simulations and theoretical analyses (e.g., Zhang (2000)) of the steady-state behavior of the accounting rate of return under conservative accounting suggest two important determinants: the long-term growth in investment and the degree of accounting conservatism. These analyses also suggest that roe approaches \(k\), but remains above it in the long-term.

Even though a decline in the excess of roe over \(k\) should cause the magnitude of abnormal earnings to fall over time, a countervailing factor is the growth in investment, which increases the base on which abnormal earnings are generated. We assume as a first approximation that the latter effect is greater than the former, and that abnormal earnings increase in perpetuity at the expected inflation rate. Since we recognize that this assumption is an approximation, we elected to err on the side of choosing too high a growth rate to ensure that our equity premium estimates are not biased downward. Also, we conduct sensitivity analyses to identify the impact on our equity premium estimates of varying the assumed growth rate within a reasonable range.

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\(^{25}\) That is, abnormal earnings persist, but show no growth. Some papers are even more conservative, and have assumed that abnormal earnings drop to zero past the horizon date.
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